

Beginning of Sept.8, 2021

Remark. S is not Gaussian. To find its distribution, one way is to notice that both S and S^2 are positive, so

$$\mathbb{P}(S \leq t) = \mathbb{P}(S^2 \leq t^2) = \mathbb{P}((n-1)S^2/\sigma^2 \leq (n-1)t^2/\sigma^2) = \mathbb{P}(\chi_{n-1}^2 \leq (n-1)t^2/\sigma^2) = \int_0^{(n-1)t^2/\sigma^2} f_{\chi_{n-1}^2}(x) dx.$$

Differentiating the above expression (along with chain rule) would give us the pdf of S .

0.1 Student's t -distribution

Recall that

$$\frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

has mean 0 and variance 1, given X_1, \dots, X_n are i.i.d. with finite mean and variance in $(0, \infty)$. Dividing both the numerator and the denominator by n gives

$$\frac{(X_1 + \dots + X_n)/n - \mu}{\sigma/\sqrt{n}} = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}.$$

Now suppose that μ, σ are unknown, and we want to find them using X_1, \dots, X_n .

It may be annoying to have two unknowns in one such equation, so sometimes we replace σ by the sample standard deviation, S , so that μ is the only free parameter, despite the fact that we don't know σ either:

$$\frac{\bar{X} - \mu}{S/\sqrt{n}}.$$

If X_1, \dots, X_n are i.i.d. *Gaussians*, then the above quotient has the **Student's t -distribution**.

Proposition: (3.7) Student's t -distribution

Let X be a standard Gaussian random variable. Let Y be a chi-squared random variable with p degrees of freedom. Assume X and Y are independent. Then

$$\frac{X}{\sqrt{Y/p}}$$

has a **student's t -distribution** with

$$f_{X/\sqrt{Y/p}}(t) = \frac{\Gamma((p+1)/2)}{\sqrt{p\pi} \Gamma(p/2)} \left(1 + \frac{t^2}{p}\right)^{-(p+1)/2}, \quad t \in \mathbb{R}.$$

Proof. First we define $Z := \sqrt{Y/p}$. It follows that for any $y > 0$,

$$\begin{aligned} f_Z(y) &= \frac{d}{dy} \mathbb{P}(Z \leq y) = \frac{d}{dy} \mathbb{P}(Y \leq y^2 p) = \frac{d}{dy} \\ &= \frac{d}{dy} \int_0^{y^2 p} \frac{x^{p/2-1} e^{-x/2}}{2^{p/2} \Gamma(p/2)} dx \\ &= 2yp \cdot p^{p/2-1} y^{p-2} e^{-y^2 p/2} \cdot \frac{1}{2^{p/2} \Gamma(p/2)} \\ &= 2y^{p-1} p^{p/2} e^{-y^2 p/2} \cdot \frac{1}{2^{p/2} \Gamma(p/2)}. \end{aligned}$$

Now we look at the CDF of X/Z :

$$\begin{aligned}\mathbb{P}(X/Z \leq t) &= \mathbb{P}(X \leq tZ) = \iint_{\substack{x \leq ty \\ y > 0}} f_{X,Z}(x, y) \, dx \, dy \\ [\text{independence}] &= \iint_{\dots} f_X(x) f_Z(y) \, dx dy.\end{aligned}\tag{\Delta}$$

Now we apply change of variable $\varphi: \mathbb{R}^2 \rightarrow \mathbb{R}^2$ by $\varphi(a, b) = (ab, a)$ and $\varphi^{-1}(x, y) = (y, x/y)$. Then

$$|\mathcal{J}(a, b)| = \text{abs value of } \begin{vmatrix} b & 1 \\ a & 0 \end{vmatrix} = |a|.$$

Then (Δ) becomes

$$\begin{aligned}\iint_{\dots} f(x, y) \, dx dy &= \iint_{\varphi^{-1}(\dots)} f(\varphi(a, b)) |\mathcal{J}(a, b)| \, da db \\ &= \iint_{\substack{b \leq t \\ a > 0}} |a| f_X(ab) f_Z(b) \, da b \\ &= \int_{b=-\infty}^{b=t} \int_{a=0}^{a=\infty} |a| f_X(ab) f_Z(b) \, da \, db.\end{aligned}\tag{\square}$$

Recall that we will eventually d/dt everything — this is exactly why we want $b = t$ as the upper limit of the outer integral: the derivative of this integral becomes

$$\begin{aligned}f_{X/Z}(t) &= \frac{d}{dt} \mathbb{P}(X/Z \leq t) \\ &= \int_{a=0}^{a=\infty} |a| f_X(at) f_Z(b) \, da \\ &= \frac{p^{p/2}}{\sqrt{2\pi}} \int_0^\infty a e^{-a^2 t^2 / 2} a^{p-1} e^{-a^2 t^2 / 2} \, dx\end{aligned}$$

□