

MATH 407 HW6

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Ch.4 Problems

4.19 The probably mass function is given by $p(0) = 1/2, p(1) = 1/10, p(2) = 1/5, p(3) = 1/10, p(3.5) = 1/10$, and 0 for all other $x \in \mathbb{R}$.

Ch.4 Theoretical Exercises

4.10

$$\begin{aligned} E[1/(X+1)] &= \sum_{k=0}^n \frac{1}{k+1} \binom{n}{k} p^k (1-p)^{n-k} \\ &= \sum_{k=0}^n \frac{n!}{(n-k)!(k+1)!} p^{k+1} (1-p)^{n-k} \\ &= \frac{1}{(n+1)p} \sum_{k=1}^{n+1} \frac{(n+1)!}{(n+1-k)!k!} [p^k (1-p)^{n+1-k}] \\ &= \frac{1}{(n+1)p} \cdot (1 - (1-p)^{n+1}). \end{aligned}$$

4.28 Since X is a geometric random variable, the sequence $\{p_n\}$ defined by $p_n := P(X = n)$ form a geometric sequence, and the relative ratio between p_n and p_m depends only on how far are they away from each other, i.e., $|m-n|$, not where m and n are exactly. Therefore the probability of $X = n+k$ given $X > n$ is the same as the probability of $X = (n+k) - n$ given $X > (n-n)$, as long as they both sense, which they indeed do (and the latter is simply “the probability of $X = k$ ”, since “ $X > n - n = 0$ ” is vacuously true).

$$\begin{aligned} P(X = n+k \mid X > n) &= P(X = n+k)/P(X > n) \\ &= p(1-p)^{n+k-1} / \sum_{i=n}^{\infty} p(1-p)^i \\ &= p(1-p)^{n+k-1} / [p(1-p)^n / p] \\ &= p(1-p)^{k-1} = P(X = k). \end{aligned}$$

4.30 Assuming the parameters are n, N, m :

$$\begin{aligned} P(X = k+1)/P(X = k) &= \left[\binom{m}{k+1} \binom{N-m}{n-k-1} \binom{N}{n} \right] / \left[\binom{m}{k} \binom{N-m}{n-k} \binom{N}{n} \right] \\ &= \left[\frac{m!}{(k+1)!(m-k-1)!} \frac{k!(m-k)!}{m!} \right] \left[\frac{(N-m)!}{(N-m-n+k+1)!(n-k-1)!} \frac{(N-m-n+k)!(n-k)!}{(N-m)!} \right] \\ &= \frac{m-k}{k+1} \cdot \frac{n-k}{N-m-(n-k-1)}. \end{aligned}$$

4.37 (a) To have an event where $X + Y$ evaluates to z_k we need some $\tilde{x} \in \{x_i\}, \tilde{y} \in \{y_i\}$ such that $\tilde{x} + \tilde{y} = z_k$. Summing all possible scenarios together we have the desired equality.

(b)

$$\begin{aligned}
 E[X + Y] &= \sum_{\tilde{z} \in \{z_k\}} [\tilde{z} \cdot P(X + Y = \tilde{z})] \\
 &= \sum_{\tilde{z} \in \{z_k\}} \left[\tilde{z} \cdot \sum_{(i,j) \in A_k} P(X = x_i, Y = y_j) \right] \\
 &= \sum_{\tilde{z} \in \{z_k\}} \sum_{(i,j) \in A_k} [\tilde{z} \cdot P(X = x_i, Y = y_j)] \\
 &= \sum_{\tilde{z} \in \{z_k\}} \sum_{(i,j) \in A_k} [(x_i + y_j) P(X = x_i, Y = y_j)].
 \end{aligned}$$

(c) Bad notation, but $\{x_i\}, \{y_j\}, \{z_k\}$ denote the possible values of X, Y , and $X + Y$. $A = \{(i, j) \mid x_i + y_j = z_k, z_k \in \{z_k\}\}$ and let $B = \{(i, j) \mid x_i \in \{x_i\}, y_j \in \{y_j\}\}$. If $(i, j) \in B$ then by definition $x_i + y_j = z_k$ for some $z_k \in \{z_k\}$ and thus $B \subset A$. If $(i, j) \in A$ then clearly $x_i \in \{x_i\}, y_j \in \{y_j\}$. Therefore $A = B$ and the two ways to write indices are equivalent.

(d)

$$P(X = x_i) = P(X = x_i, Y \in \mathbb{R}) = P(X = x_i, Y \in \{y_j\}) = \sum_{y_j \in \{y_j\}} P(X = x_i, Y = y_j)$$

where the last step is true since the events are pairwise disjoint when we let y_j vary.

(e) This I believe has already been shown in class twice.

$$\begin{aligned}
 E[X + Y] &= \sum_{x_i \in \{x_i\}} \sum_{y_j \in \{y_j\}} (x_i + y_j) P(X = x_i, Y = y_j) \\
 &= \sum_{x_i \in \{x_i\}} \sum_{y_j \in \{y_j\}} x_i \cdot P(X = x_i, Y = y_j) + \sum_{x_i \in \{x_i\}} \sum_{y_j \in \{y_j\}} y_j \cdot P(X = x_i, Y = y_j) \\
 &= \sum_{x_i \in \{x_i\}} x_i \sum_{y_j \in \{y_j\}} P(X = x_i, Y = y_j) + \sum_{y_j \in \{y_j\}} y_j \sum_{x_i \in \{x_i\}} P(X = x_i, Y = y_j) \\
 &= \sum_{x_i \in \{x_i\}} x_i P(X = x_i) + \sum_{y_j \in \{y_j\}} y_j \sum_{x_i \in \{x_i\}} P(X = x_i, Y = y_j) \\
 &= \sum_{x_i \in \{x_i\}} x_i P(X = x_i) + \sum_{y_j \in \{y_j\}} y_j P(Y = y_j) \\
 &= E[X] + E[Y].
 \end{aligned}$$