

**Theorem**

Let  $\{X_{i,j} : i \leq n, j \leq m(i)\}$  be independent. Then  $f(X_{i,1}, \dots, X_{i,m(i)})$ ,  $i \leq n$ , are independent random variables.

*Proof.* Let  $\mathcal{F}_{i,j} = \sigma(X_{i,j})$  and  $\mathcal{B}_i := \sigma(\mathcal{F}_{i,1}, \dots, \mathcal{F}_{i,m(i)})$ . By the previous theorem each  $\mathcal{B}_i$ 's are independent. Each  $f_i$  is  $\mathcal{B}_i$ -measurable so the random variables  $f_i$  are independent.  $\square$

Fubini theorem says for  $f(x, y)$  on  $\Omega \times \Omega_2$ ,

$$\int f \, d(\mu_1 \times \mu_2) = \int_{\Omega_1} \int_{\Omega_2} f \, d\mu_2 \, d\mu_1$$

provided  $f \geq 0$  or  $f$  is integrable (i.e.,  $\int |f| \, d(\mu_1 \times \mu_2) < \infty$ ). (Here since  $\mu_i$ 's are probability measures they are assumed to be  $\sigma$ -finite.) For random variables:

**Theorem: D2.1.12**

Let  $X, Y$  be independent with distributions  $\mu_X$  and  $\mu_Y$  on  $\mathbb{R}$ . Let  $h : \mathbb{R}^2 \rightarrow \mathbb{R}$  satisfy either  $h \geq 0$  or  $\mathbb{E}|h(X, Y)| < \infty$ . Then

$$\mathbb{E}h(X, Y) = \iint_{\mathbb{R}^2} h(X, Y) \, d\mu_X(dx) \, d\mu_Y(dy) \quad (*)$$

and the order of integration does not matter.

In particular, for  $h(x, y) = f(x)g(y)$  with either  $f, g \geq 0$  or  $\mathbb{E}|f(X)|, \mathbb{E}|g(Y)| < \infty$ , we have

$$\mathbb{E}[f(X)g(Y)] = \mathbb{E}f(X)\mathbb{E}g(Y), \quad (**)$$

i.e., independence  $\Rightarrow$  (product of  $\mathbb{E} = \mathbb{E}$  of product).

*Proof.* (\*) follows from Fubini since the distribution of  $(X, Y)$  is  $\mu_X \times \mu_Y$  by independence.

For (\*\*),

$$\begin{aligned} \mathbb{E}[f(X)g(Y)] &= \iint_{\mathbb{R}^2} f(x)g(y) \, \mu_X(dx) \, \mu_Y(dy) \\ &= \int_{\mathbb{R}} g(y) \int_{\mathbb{R}} f(x) \, \mu_X(dx) \, \mu_Y(dy) \\ &= \left( \int_{\mathbb{R}} f(x) \, \mu_X(dx) \right) \int_{\mathbb{R}} g(y) \, \mu_Y(dy) = \mathbb{E}f(X)\mathbb{E}g(Y). \end{aligned} \quad \square$$

By induction, we may generalize the above result into any finite number of random variables. That is, for independent  $X_1, \dots, X_n$  with  $\mathbb{E}|\prod X_i| < \infty$ , we have  $\mathbb{E}[\prod X_i] = \prod \mathbb{E}X_i$ .

**Sums of Independent Random Variables**

Let  $X, Y$  be independent with distribution functions  $F$  and  $G$ . The d.f. of  $X + Y$  is the **convolution**

$$H(z) = \mathbb{P}(X + Y \leq z) = \int_{-\infty}^{\infty} F(z - y) \, d(G(y)) =: (F * G)(z).$$

To see this, we apply Fubini to  $1_{x+y \leq z}$ :

$$H(z) = \mathbb{E}1_{\{X+Y \leq z\}} = \iint 1_{\{x \leq z-y\}} dF(x) dG(y) = \int F(z-y) dG(y).$$

Note by doing  $\{y \leq x - z\}$  first we obtain see that  $F * G \equiv G * F$ .

**Example.** Let  $X$  be uniform on  $[0, 2]$  and  $Y$  is exponential with parameter  $\lambda$ . That is,  $X$  has  $1/2$  on  $[0, 2]$  and  $Y$  has  $\lambda e^{-\lambda y}$  on  $[0, \infty)$ . The distribution function of  $Y$  is  $1 - e^{-\lambda y}$  for  $y \geq 0$ . Then

$$H(z) = \int_{-\infty}^{\infty} \underbrace{(1 - e^{-\lambda(z-y)})}_{F(z-y)} 1_{\{z-y \geq 0\}} \underbrace{\frac{1}{2} 1_{[0,2]}}_{dG(y)} dy.$$

That is,

$$H(z) = \begin{cases} 0 & z < 0 \\ \int_0^z (1 - e^{-\lambda z} e^{\lambda y})/2 dy & 0 \leq z \leq 2 \\ \int_0^2 (1 - e^{-\lambda z} e^{\lambda y})/2 dy & z > 2 \end{cases} = \begin{cases} 0 & z < 0 \\ \frac{z}{2} - \frac{1 - e^{-\lambda z}}{2\lambda} & 0 \leq z \leq 2 \\ 1 - e^{-\lambda z} \frac{e^{2\lambda} - 1}{2\lambda} & z > 2. \end{cases}$$

Let  $\mu_n$  be a probability measure on  $(\mathbb{R}^n, \mathcal{R}^n)$ . We can make a random vector with distribution  $\mu_n$ : we take  $(\Omega, \mathcal{F}, \mathbb{P}) = (\mathbb{R}^n, \mathcal{R}^n, \mu_n)$  and  $(X_1, \dots, X_n)$  to be identity.

## Infinite Sequence of Random Variables

We say **finite-dimensional** distributions of  $\{X_n, n \geq 1\}$  are all distributions of form  $\{X_i, i \in I\}$  for  $I \subset \mathbb{N}$  finite. By using marginals is sufficient to consider  $I = \{1, \dots, n\}$ . Suppose we are given  $\mu_n$  on  $(\mathbb{R}^n, \mathcal{R}^n)$  for every  $n$ .

**Question:** is there a  $\mathbb{P}$  on  $(\mathbb{R}^{\mathbb{N}}, \mathcal{R}^{\mathbb{N}})$  with distribution  $\mu_n$  for the first  $n$  coordinates? That is,

$$\mathbb{P}(A_1 \times \dots \times A_n \times \mathbb{R} \times \dots) = \mu_n(A_1 \times \dots \times A_n)?$$

(Well of course no, since if  $n > m$ ,  $\mu_n$  determines what  $\mu_m$  would be.) What if this consistency is satisfied?

### Theorem: Kolmogorov Extension Theorem

Let  $\mu_n$  be a p.m. on  $(\mathbb{R}^n, \mathcal{R}^n)$  for all  $n$ . Suppose consistency holds among the  $\mu_n$ 's, i.e.,

$$\mu_{n+1}(A \times \mathbb{R}) = \mu_n(A) \quad \text{for all } A = \prod_{i=1}^n (a_i, b_i] \text{ and } n \geq 1.$$

(In reality the above choice of  $A$  can be anything in  $\mathcal{R}^n$ ; we just picked the most canonical one.) Then there exists a unique  $\mathbb{P}$  on  $(\mathbb{R}^{\mathbb{N}}, \mathcal{R}^{\mathbb{N}})$  with

$$\mathbb{P}\left(\prod_{i=1}^n (a_i, b_i] \times \mathbb{R} \times \mathbb{R} \times \dots\right) = \mu_n\left(\prod_{i=1}^n (a_i, b_i]\right).$$

(Sets of form  $A \times \mathbb{R} \times \mathbb{R} \times \dots$  with  $A \in \mathcal{R}^n$  is a **cylinder set** in  $\mathbb{R}^{\mathbb{N}}$ . They form a  $\sigma$ -field.)

## 0.1 Weak Laws of Large Numbers

Some recaps:

- We say  $X_n \rightarrow X$  **in probability** (i.e. in measure) if  $\mathbb{P}(|X_n - X| > \epsilon) \rightarrow 0$  as  $n \rightarrow \infty$ .
- We say  $X_n \rightarrow X$  **in  $L^p$**  (where  $p > 0$ ) if  $\mathbb{E}(|X_n - X|^p) \rightarrow 0$  as  $n \rightarrow \infty$ .
  - For  $p \geq 1$ , this is equivalent to  $\|X_n - X\|_p \rightarrow 0$ .
  - For  $p < 1$  this does not hold as such  $\|\cdot\|_p$  does not define a norm.
  - In principle the  $X_n$  can have infinite  $p^{\text{th}}$  moment but the definition still makes sense.